

#### THE UNIVERSITY OF ZAMBIA

#### SOCIOLOGY

A NOTE ON THE "L" TEST AND THE "L" " TEST AS USEFUL ALTERNATIVES TO THE X<sup>2</sup> TEST IN THE ANALYSIS OF SMALL-SAMPLE DATA

Wim M.J. van Binsbergen

December 1972

# Contents:

-Íntroduction

-Disadvantages of the  $\chi^2$  test

-The 1 test

-The 1' test

-References

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# Introduction

In social statistics, we may use the  $\chi^2$  test if we wish to analyze (typically in data derived from a small sample, and about the underlying distribution of which we do not want to make assumptions — so that we cannot use a parametric test):

- 1. Whether there exists a statistically significant difference between two or more sets of tabulated, empirical (stochastical) data. Typically:

 					Service of the servic
	X	Y	· X+Y		
A	p	q ·	p+q		table 1
В	r	S	r+s		
A+B	p+r	q+s	p+q+r+s	•	8 #

Example: We take an aselect sample of 100 UNZA students; the variables X and Y stand for, resp., male and female; A and B stand for, resp. baptized and non-baptized in any Christian church. Our table could have the following form:

	mole	female	total	
baptized	70	8	78	ret re
non-bapt.	19	3	22	table 2
total	89	11	100	

The  $\chi^2$  test now yields a particular statistic (in this case:  $\chi^2$ = 0.20), associated with a particular number of degrees of freedom (df = (a-1).(b-1) = 1, where a is number of row entries and b is number of column entries.) Consulting a  $\chi^2$  table we

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В	r	S	r+s		
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The  $\frac{2}{X}$  test now yields a particular statistic (in this case:  $\frac{2}{X}$ = 0.20), associated with a particular number of degrees of freedom (df = (a-1).(b-1) = 1, where a is number of row entries and b is number of column entries.) Consulting a  $\frac{2}{X}$  table we

very suitable for the research problems arising from small sample research in the social sciences (Binsbergen 1970). These tests were called by Spitz: 1 test (for two or more stochastic distributions) and 1' test (for one stochastic distribution as compared to a standard expectation); the name derives from the extensive use of 1 ogarithms that is made in these tests. Once one is accustomed to the use of (natural) logarithms, calculations are extremely simple; none of the disadvantages of the  $_{\rm X}^{2}$  occur (there is no correction for continuity meither a minimal requirement per cell); and, once the statistic is calculated, it can be interpreted with a normal  $_{\rm X}^{2}$  table, observing the same rules about degrees of freedom as apply to  $_{\rm X}^{2}$ . (1)

I will-not discuss the mathematical background of the land l' tests in relation to the  $\chi^2$  test (cf. Mood 1950; Woolf 1957; Spitz 1961), but shall only demonstrate the application of these tests on the basis of the (fictive) data used previously (tables 2 and 4).

### The 1 test.

The general formula for the computation of 1 is:

$$1 = 2 / \ln n + \sum_{i=1}^{a} \sum_{j=1}^{b} f_{ij} \cdot \ln f_{ij} - \sum_{i=1}^{a} n_{i} \cdot \ln n_{i} - \sum_{j=1}^{b} f_{j} \cdot \ln f_{j} / 7$$

(where:

n = sample size

ln = natural logarith (basis e)

a = number of row entries

i = rank number of a particular row entry (1) i a)

b = number of column entries

j = rank number of a particular column entry(1) j b)

f<sub>ij</sub> = frequency in a particular, identified cell

n<sub>i</sub> = bottom total of a particular column

f<sub>i</sub> = right-hand total of a particular row)

This may look extremely complicated, but things become clearer if we rewrite table 1 accordingly (p.t.o.):

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<sup>(1)</sup> Thus, for the 1 test, df = (a-1).(b-1); for the 1' test, df = a - 1.

	Х	Y	X+Y	
A	f <sub>ll</sub>	<sup>f</sup> 21	fl	
B + ( sto + 7	<sup>f</sup> 12	f <sub>22</sub>	f <sub>2</sub>	table 5
A+B	n <sub>1</sub>	n <sub>2</sub>	'n	

Substituting now the empirical values as shown in table 2, the statistic is computed as follows:

$$1 = 2\sqrt{(100.\ln 100) + (70.\ln 70 + 8.\ln 8 + 19.\ln 19 + 3.\ln 3) - (89.\ln 89 + 11.\ln 11) - (78.\ln 78 + 22.\ln 22)} = 0.198.$$

The various values of a.ln a can be calculated on the basis of a table of natural logarithms, but of course it saves much time if one uses a table of a.ln a straight-away; such a table is given by Spitz (1961; 1965).

So the standard procedure for the 1 test, no matter the number of row and column entries in our table, is:

- 1. add the values of a.ln a for, successively, n (=
   sample size) and all cell frequencies;
- 2. substract from this sum the values of a.ln a for all bottom and all right-hand totals;
- 3. multiply the result by two.

As you see, the statistic 'thus obtained has exactly the same value as  $\chi^2$  would have had; differences that may show are due to rounding-off.

# The l' test.

The general formula for the computation of l' is:

$$1' = 2 \sum_{i=1}^{a} f_{i} \cdot \ln \left( \frac{f_{i}}{e_{i}} \right)$$

(where:

f = frequency observed in the i-th column entry e = frequency expected in the i-th column entry

ln = natural logarithm

i = rank number of a particular column entry(1) i)a)
a = number of column entries.)

Substituting the empirical and expected values as shown in table 4, we get:

$$1' = 2 \left( 70.\ln \frac{70}{(40/100.89)} + 19.\ln \frac{19}{(60/100).89} \right) =$$

$$2 \left( 70.\ln \frac{700}{356} + \frac{190}{534} \right) =$$
(Note that  $\ln \frac{a}{b} = \ln a - \ln b$ )

$$2\sqrt{70.(\ln 700 - \ln 356) + 19.(\ln 190 - \ln 534)} = 55.4.$$

Jr. 39282 This is again the same result as  $_{
m X}^{
m 2}$  would have given. For the calculation of l' a table of natural logarithms is obviously indispensable.

The standard procedure for the computation of 1' (no matter the number of column entries, i.e. of categories, in our data) is:

- 1. Calculate the expected value as a fraction of the total sample size:
- 2. per column entry, divide the observed value by the expected value, and simplify-without writing the fraction as a decimal fraction (this because it is tedious to find the natural logarithms of decimal fractions);
- 3. per column entry, subtract the natural logarithm of the expected value from the natural logarithm of the observed value - and multiply the result by the observed value;
- 4. add these products for each column entry;
- 5. multiply the sum by 2.

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